



**STUDY OF A NEW POWER DISTRIBUTION AS A FAILURE MODEL : ITS STATISTICAL
AND MATHEMATICAL PROPERTIES**

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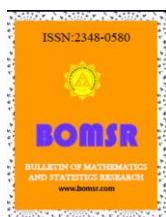
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ABSTRACT

A finite range probability distribution has been developed and some of its statistical and structural properties have been discussed. Goodness of fit has been successfully tested on the Transmitter tube failure data of Davis (1952).
Keywords: Probability distribution, statistical properties, structural properties, Goodness of fit.

Introduction

Introducing new probability distributions have been an interesting field of research for last many decades. Weibull (1951), Folks and Chikara (1978), Mukherji and Islam (1983), Siddiqui et al (1992, 1994, 1995, 2016), Cha (2007), Madal and Yakov (2010) and Shradha et al (2017) and Syed A.A. et al (2016).

The proposed probability distribution is also useful in many real life data so it can be added in family of probability distributions.

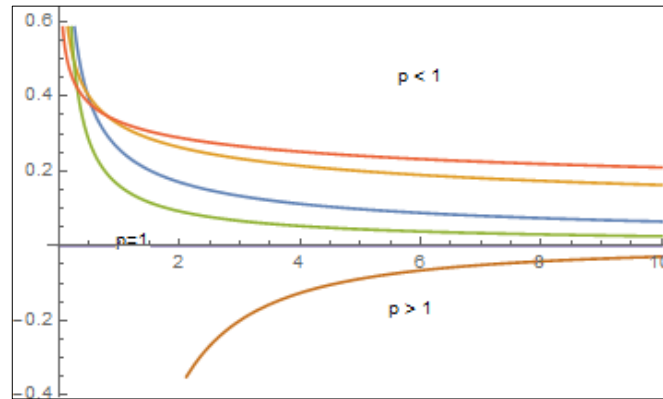
1 Proposed distribution

The probability density function of the proposed distribution is:

$$f(x) = (1 - p)x^{-p} \theta^{p-1} \quad (1)$$

here p is the shape parameter and θ is the scale parameter.

$$0 < x < \theta; \theta > 0; 0 < p < 1$$



Graph(1): pdf

Conditions: $p < 1 \rightarrow$ Shape is reverse J shape

$p = 1 \rightarrow$ Shape is straight line

$p > 1 \rightarrow$ shape is inverse bath tub

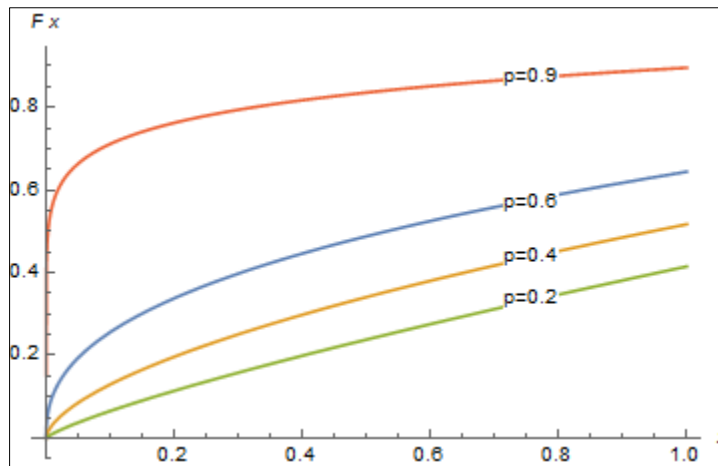
For lesser value of parameter p proposed distribution is a growth function and for higher value of p it is treated as a decay function. Therefore it can be used for profit and loss both in financial data.

The cumulative distribution is defined as:

$$F(x) = \theta^{p-1} x^{1-p} \tag{2}$$

where,

$$0 < x < \theta; \theta > 0; 0 < p < 1$$



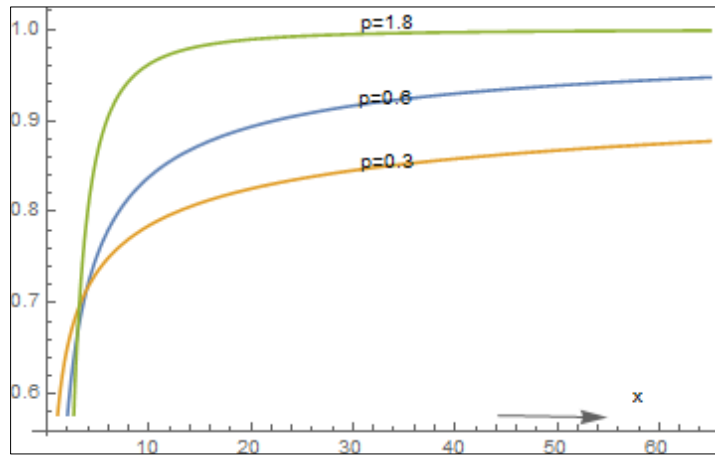
Graph(2): CDF

1.1 Reliability Function

Reliability function of proposed model is

$$R(x) = 1 - F(x) = 1 - \theta^{p-1} x^{1-p} \tag{3}$$

where θ and p are scale and shape parameters respectively. at $t = 0$; eqn 3 gives $R(0) = 1$ and as time t increases reliability decreases.



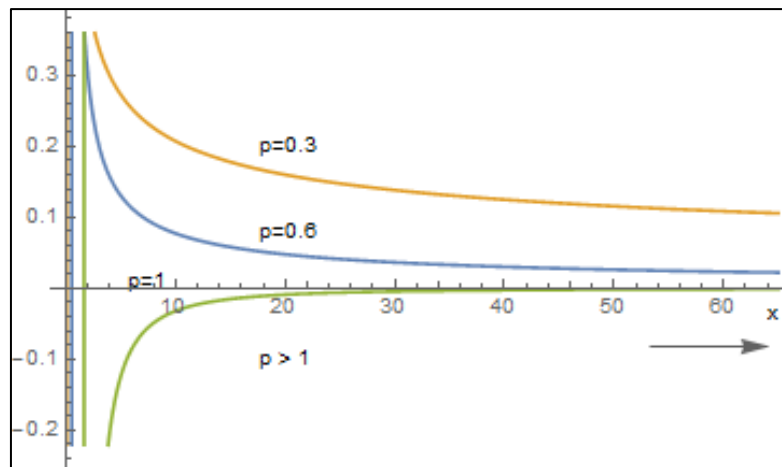
Graph(3) : Reliability function

1.2 Hazard Rate Function

Hazard rate or Failure rate shows the failures of engineered system per unit of time. For the new distribution it is shown by

$$\frac{f(x)}{1 - F(x)} = \frac{(1 - p)x^{-p}\theta^{p-1}}{1 - \theta^{p-1}x^{1-p}}$$

$$= \frac{(1 - p)x^{-p}}{\theta^{1-p}x^{1-p}} \dots \dots \dots (4)$$



Graph(4) : Hazard Rate Function

Graph shows the given distribution has only reverse J-shaped [DFR (Burn in) < CFR (Useful life) Hazard rate it means hazard rate is monoton- ically decreasing. The reverse-J hazard rate is a better tool over hazard rate for the characterization of any failure mode. D De- sai, V. Maariapan et.al [2010].

1.3 Moment Generating function (mgf)

Moment generating function is an important statistical property of a probability distribution function as it provides the basis of an alternative route to analytical results compared to directly with pdf or cdf .

If any random variable x has new proposed distribution, then the moment generating (mgf) will be:

$$\begin{aligned}
 M_x(t) &= E(e^{tx}) \\
 &= \int_0^\theta (1-p)\theta^{p-1}x^{-p} dx \\
 &= (1-p) \sum_{r=0}^{\infty} \frac{t^r \theta^r}{r! [(r+1)-p]} \\
 M_x(t) &= (1-p) \sum_{r=0}^{\infty} \frac{t^r \theta^r}{r! [(r+1)-p]} \quad (5)
 \end{aligned}$$

which is the desired moment generating function.

1.4 Mean

The Mean μ_r a continuous random variable of proposed probability density function pdf is

$$\begin{aligned}
 \mu'_r &= E(X) = \int_0^\theta x f(x) dx \\
 \mu'_r &= E(X) = \int_0^\theta (1-p)\theta^{p-1}x^{-p} dx
 \end{aligned}$$

where, $0 < x < \theta$; $\theta > 0$; $0 < p < 1$

$$\mu'_1 = \frac{1-p}{(r+1)-p} \theta^r \quad (6)$$

Using above relation(6), we get first and second moment as:

$$\mu'_1 = E(X) = \frac{1-p}{2-p} \theta \quad (7)$$

$$\mu'_1 = E(X^2) = \frac{1-p}{3-p} \theta^2 \quad (8)$$

μ'_1 refers to the mean of the proposed distribution.

1.5 Variance

Variance the measure of spread can be find by below:

$$\begin{aligned}
 V_{(x)} &= \mu'_2 - (\mu'_1)^2 \\
 &= E(X^2) - [E(X)]^2 \\
 &= \left[\frac{1-p}{3-p} - \left(\frac{1-p}{2-p} \right)^2 \right] \theta^2 \\
 V_{(x)} &= \frac{(1-p)\theta^2}{(3-p)(2-p)^2} \quad (9)
 \end{aligned}$$

above equation (9) shows the variance of the proposed distribution.

1.6 Median

$$\int_0^{Me} (1-p)x^{-p}\theta^{p-1}dx = \frac{1}{2}$$

$$\begin{aligned} (x^{1-p} \theta^{p-1})_0^{Me} &= \frac{1}{2} \\ (Me)^{1-p} \theta^{p-1} &= \frac{1}{2} \\ (Me)^{1-p} &= \frac{\frac{1}{2}}{\theta^{p-1}} \\ \ln(Me)^{1-p} &= \ln \frac{1}{2} - \ln \theta^{p-1} \\ Me &= \frac{\ln \frac{1}{2} - \ln \theta^{p-1}}{1-p} \quad (10) \end{aligned}$$

1.7 Mode

For Mode on differentiating $f(x)$ with respect to x , we get

$$\begin{aligned} f'(x) &= (1-p)(-p)x^{-p-1}\theta^{p-1} \\ f''(x) &= (1-p)(-p-1)x^{-p-2}\theta^{p-1} \\ f''(x) &= -(p+2)(p+1)p(1-p)x^{-p-3}\theta^{p-1} > 0 \end{aligned}$$

it indicates that mode for the new distribution does not exist.

1.8 Forgetfulness Property

The distribution is forgetful or has no memory. Means if a unit survived for the t hours, then the probability of its surviving an additional h hours is exactly same

$$\begin{aligned} P(X \geq t+h | X \geq t) &= \frac{\int_t^{t+h} (1-p)x^{-p}\theta^{p-1} dt}{\int_0^t (1-p)x^{-p}\theta^{p-1} dt} \\ &= \left(1 + \frac{h}{t}\right)^{1-p} - 1 \end{aligned}$$

which is free from x , which shows the forgetfulness property of the proposed distribution.

2 Estimation of parameters

2.1 Maximum Likelihood estimation (mle)

Let x_1, x_2, \dots, x_n be a random sample of size n , Likelihood estimation of the parameters can be written as

$$\begin{aligned} L(p) &= \prod_{t=1}^n (1-p)x_i^{-p}\theta^{p-1} \\ &= (1-p)^n \theta^{n(p-1)} \prod_{t=1}^n x^{-p} \\ \frac{\partial \ln(p)}{\partial p} &= 0 \Leftrightarrow \frac{-n}{1-p} + n \ln \theta - \sum_{t=1}^n \ln(x) \\ \hat{p} &= \frac{n}{n \ln \theta - \sum \ln x} \quad (11) \end{aligned}$$

above equation(11) shows estimation for the shape parameter p .

$$\hat{\theta} = t_n = \text{Max} (t_1, t_2, \dots, t_n) \quad (12)$$

above equation(12) shows estimation for the shape parameter θ .

2.2 Moments method

$$\begin{aligned} \text{Mean} &= \frac{1-p}{2-p} \theta = GM \\ (1-p)\theta &= (2-p)GM \\ p(GM - \theta) &= 2GM - 1 \\ \hat{p} &= \frac{2GM-1}{GM-\theta} \quad (13) \end{aligned}$$

3 Order statistics

Here the density $f_{i:n}$ of the i th order statistics, for $i = 1, 2, \dots, n$ from independent identically distributed random variable X_1, X_2, \dots, X_n is given by

$$\begin{aligned} f_{i:n}(x) &= \frac{f(x)}{B(i, n-i+1)} = F(x)^{i-1} [1-F(x)]^{n-i} \quad (14) \\ f_{i:n}(x) &= \frac{(-p)x^{-p}\theta^{p-1}}{B(i, n-i+1)} \sum_{k=0}^{n-1} (-1)^k \binom{n-1}{k} [x^{1-p}\theta^{p-1}]^{(1+k)-1} \end{aligned}$$

where $0 < F(x) < 1$ for $x > 0$

by above equation we can express K_{th} ordinary moment of the i th order statistics $X_{i:n}$ as a linear combination of the K_{th} moment of the new distribution with different shape parameter. Therefore, the measure of skewness and kurtosis of the distribution of $X_{i:n}$ can be calculated.

The L-moments are analogous to the ordinary moments but can be estimated by linear combinations of order statistics.

$$\lambda_1 = E[X_{1:1}] \quad (15)$$

$$\lambda_2 = \frac{1}{2} E[X_{2:2} - X_{1:2}] \quad (16)$$

$$\lambda_3 = \frac{1}{3} E[X_{3:3} - 2X_{2:3} + X_{1:3}] \quad (17)$$

$$\lambda_4 = \frac{1}{4} E[X_{4:4} - 3X_{3:4} + 3X_{2:4} - X_{1:4}] \quad (18)$$

4 Fisher information Matrix of new distribution

The Fisher information is that a random variable 'X' contains about the parameter θ is given by

$$I(\theta) = E\left[\frac{\partial}{\partial x \log f(x; \theta)}\right]^2 \quad (19)$$

Now if $\log f(x; \theta)$ is twice differentiable w.r.t. θ under certain regularity conditions, Fisher's information is given by;

$$I(\theta) = E_{\theta}[\partial^2/\partial\theta^2 \log\{f(x; \theta, p)\}] \quad (20)$$

Where

$$f(x; \theta, p) = (1-p)x^{-p}\theta^{p-1}$$

$$\log\{f(x; \theta, p)\} = \log(1-p) - p\log x + (p-1)\log\theta \quad (21)$$

differentiating w.r.t. p , θ and taking Expectations on both sides of the those equations we get following equations:

$$-E\left[\frac{\partial \log f(x; \theta, p)}{\partial p^2}\right] = \frac{1}{(1-p)^2} \quad (22)$$

$$-E\left[\frac{\partial^2 \log f(x; \theta, p)}{\partial p \partial \theta}\right] = \frac{-1}{\theta} \quad (23)$$

$$-E\left[\frac{\partial^2 \log f(x; \theta, p)}{\partial \theta \partial p}\right] = \frac{-1}{\theta} \quad (24)$$

$$-E\left[\frac{\partial^2 \log f(x; \theta, p)}{\partial \theta^2}\right] = \frac{(1-p)}{\theta^2} \quad (25)$$

Now we have

$$I_{(1,1)} = \frac{1}{(1-p)^2}$$

$$I_{(1,1)} = \frac{-1}{\theta}$$

$$I_{(1,2)} = \frac{-1}{\theta}$$

$$I_{(1,1)} = \frac{p-1}{\theta^2}$$

Above equation is the Fisher's information Matrix of the new distribution.

5 Entropies

5.1 R'enyi and q-entropy

The entropy of a random variable X is a measure of the uncertain variations in the system. The R'enyi entropy is defined by

$$I_R(\delta) = \frac{1}{(1-\delta)} \log[I(\delta)]$$

Where

$$I(\delta) = \int_R f^\delta(x) dx \text{ for } \delta > 0 \text{ and } \delta \neq 1$$

$$f^\delta = (1-p)^\delta (x^{-p})^\delta (\theta^{p-1})^\delta$$

$$I(\delta) = \int_R f^\delta = (\theta^{p-1})^\delta (x^{1-p})^\delta$$

Hence R'enyi entropy reduces to

$$I_R(\delta) = \frac{1}{(1-\delta)} \log[(\theta^{p-1})^\delta (x^{1-p})^\delta]$$

5.2 q-entropy

q – entropy say $H_q(f)$ is defined by

$$H_q(f) = \frac{1}{(q-1)} \log[1 - I_q(f)]$$

Where

$$I_q(f) = \int_R f^q(x) dx \text{ for } q > 0 \text{ and } q \neq 1$$

$$H_q(f) = \frac{1}{(q-1)} \log[1 - (\theta^{p-1})^q (x^{1-p})^q]$$

5.3 Shannon's entropy

$$f(x) = (1-p)x^{-p}\theta^{p-1}$$

$$\log L^*(x, \theta, p) = n \log(1-p) + n(p-1) \log \theta - p \sum_{i=1}^n \log x_i$$

$$l(x, \theta, p) = n [\log(1-p) + (p-1) \log \theta] - p \sum_{i=1}^n \log x_i$$

$$\frac{l(x, \theta, p)}{n} = \log(1-p) + (p-1) \log \theta - \overline{p \log x_i}$$

The Shannon's entropy of new distribution will be

$$\hat{H} = -[\log(1-p) + (p-1) \log \theta - \overline{p \log x_i}]$$

$$\hat{H} = \frac{-l(x; \theta, p)}{n} \quad (26)$$

above equation shows the Shannon's entropy of new distribution.

6 Goodness of Fit

To check the goodness of fit of the proposed distribution a data set of Radar set component failure is taken for a Transmitter Tube failure (D.J.Davis-1952).

1-Transmitter Tube

$t(hr.)$	Observed freq.	Expected freq.	$(o - e)^2/e$
0-25	109	105	0.1524
25-50	42	30	4.88
50-75	17	21	0.7619
75-100	7	17	5.8823
100-up	13	15	0.2666
	188	188.0	11.9432

$$d.f. = 4; \chi^2 = 11.9432; p = 0.6$$

$\chi_{cal}^2 = 11.9432$ and $\chi_{tab}^2 = 13.28$ at level of significance. it shows that the proposed data fits well to this data of failure.

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